

Banerjee, S.N. (2022). Estimation of high blood lead levels among children in Georgia: An application of Bayesian analysis. *Journal of Environmental Health*, 85(3), 8–15.

Corresponding Author: Shailendra N. Banerjee, Mathematical Statistician, Division of Environmental Health Science and Practice, National Center for Environmental Health, Centers for Disease Control and Prevention, Atlanta, GA 30341. Email: snb1@cdc.gov.

Note. This supplemental file was submitted by the authors along with the respective peer-reviewed article and has been posted online due to space limitations at <https://www.neha.org/jeh/supplemental>. The Journal of Environmental Health did not copy (d)-5.4neo lth We used “z” to represent the number of children aged <6 years with BLLs of 5–9 µg/dL in a county

distribution with parameter “ ” and “m” as the number of children age <6 years who were tested

for BLL (equation 1):

$$p(z/) = e$$

$$^{-m \cdot } (m \cdot)^z / z! \tag{1}$$

where is the rate of children with BLLs of 5–9 µg/dL

i.e., = (children with BLLs of 5–9 µg/dL) / (children tested for BLL)

If follows a gamma (,) prior

$$\text{i.e., } p() = e^{- ()} \cdot^{-1} / () \tag{2}$$

where > 0, then posterior distribution of is given by

$$p(/z) = p(z/) \times p() / p(z)$$

i.e., p()

If we assume that the prior information about parameter λ , the rate of children with BLLs of 5-9 $\mu\text{g/dL}$, can be obtained from a small group of counties in Georgia, who we believe has the same rate (λ) of 5-9 $\mu\text{g/dL}$ BLLs among children aged <6 years, then the posterior for λ can be estimated from equation (3).

We suppose z_j is the number of children aged <6 years with BLLs of 5–9 $\mu\text{g/dL}$ among x_j children from county “j”. Then assuming it follows a Poisson distribution, we have the following:

$$p(z_j/\lambda) = e^{-(\lambda x_j)} (\lambda x_j)^{z_j} / z_j! \quad (4)$$

where λ is the same as defined earlier.

So, the likelihood function for n counties with the same parameter λ is given as follows:

$$L(\lambda / z_j) = e^{-(\lambda \sum x_j)} (\lambda \sum x_j)^{z_1 + z_2 + \dots + z_n} / z_1! z_2! \dots z_n! \quad (5)$$

So,

$$L(\lambda / z_j) = e^{-(\lambda \sum x_j)} (\lambda \sum x_j)^{z_j} \quad (6)$$

If for all these n counties we assume that λ follows a non-informative prior $1/\lambda$, i.e., $p(\lambda) = 1/\lambda$, then from equation (6), the posterior distribution of λ is given by the following:

$$p(\lambda / z_j) = e^{-(\lambda \sum x_j)} (\lambda \sum x_j)^{z_j} \cdot 1/\lambda$$

i.e., $p(\lambda / z_j) = e^{-(\lambda \sum x_j)} (\lambda \sum x_j)^{z_j-1}$ (7)

This is a gamma (α, β), where

$$\alpha = \sum z_j \text{ and } \beta = \sum x_j$$

county. We then can use known θ and $p(z|\theta)$ from equation (8) in equations (2) and (3) to evaluate the prior and posterior distributions of the parameter θ in the targeted county.

The joint distribution of data z and the parameter θ are given by the following:

$$p(z, \theta) = p(\theta) \times p(z|\theta), \text{ and also}$$

$$p(z, \theta) = p(z) \times p(\theta|z)$$

Thus, $p(z) \times p(\theta|z) = p(\theta) \times p(z|\theta)$, giving

$$p(z) = p(\theta) \times p(z|\theta) / p(\theta|z) \tag{9}$$

Here, $p(\theta)$ and $p(\theta|z)$ are the known prior and posterior distributions, respectively, of the parameter θ . Thus, $p(\theta)$